

Diversity and Coevolutionary Dynamics in High-Dimensional Phenotype Spaces*

Michael Doebeli^{1,†} and Iaroslav Ispolatov²

1. Departments of Zoology and Mathematics, University of British Columbia, 6270 University Boulevard, Vancouver, British Columbia V6T 1Z4, Canada; 2. Departamento de Física, Universidad de Santiago de Chile, Casilla 302, Correo 2, Santiago, Chile

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ABSTRACT: We study macroevolutionary dynamics by extending microevolutionary competition models to long timescales. It has been shown that for a general class of competition models, gradual evolutionary change in continuous phenotypes (evolutionary dynamics) can be nonstationary and even chaotic when the dimension of the phenotype space in which the evolutionary dynamics unfold is high. It has also been shown that evolutionary diversification can occur along nonequilibrium trajectories in phenotype space. We combine these lines of thinking by studying long-term coevolutionary dynamics of emerging lineages in multidimensional phenotype spaces. We use a statistical approach to investigate the evolutionary dynamics of many different systems. We find (1) that, for a given dimension of phenotype space, the coevolutionary dynamics tend to be fast and nonstationary for an intermediate number of coexisting lineages but tend to stabilize as the evolving communities reach a saturation level of diversity and (2) that the amount of diversity at the saturation level increases rapidly (exponentially) with the dimension of phenotype space. These results have implications for theoretical perspectives on major macroevolutionary patterns such as adaptive radiation, long-term temporal patterns of phenotypic changes, and the evolution of diversity.

Keywords: long-term evolution, diversity and stability, adaptive radiation.

Introduction

One of the fundamental problems in evolutionary biology is to understand how microevolutionary processes generate macroevolutionary patterns. In particular, the emergence of both macroevolutionary changes in the speed of evolution (Simpson 1944; Gould and Eldredge 1977) and macroevolutionary changes in patterns of species diversity (Rosenzweig 1995; Schluter 2000) has long been of great interest. For example, Uyeda et al. (2011) have recently proposed

that, over macroevolutionary timescales, relatively short, intermittent bursts of high rates of evolutionary change should alternate with long periods of bounded phenotypic fluctuations. Also, there is much discussion about whether species diversity saturates over evolutionary time in a given environment (Harmon and Harrison 2015; Rabosky and Hurlbert 2015; Rosenzweig 1995). Phylogenetic analysis has been used to shed light on these questions (Pennell and Harmon 2013), but mechanistic models in which short-term ecological interactions are extrapolated to yield long-term patterns of diversity and evolutionary change have only recently been developed. Most of these models have been used to study the long-term evolution of diversity by analyzing processes of community assembly emerging from short-term ecological dynamics (Loeuille and Loreau 2005, 2009; Ito and Dieckmann 2007; Allhoff et al. 2015; Gascuel et al. 2015; Rosindell et al. 2015). In particular, these papers have mainly focused on how diversity changes over time but not on how the nature of the coevolutionary dynamics of a given set of coexisting species changes as the diversity changes. In fact, in all these models, the evolutionary dynamics for a fixed amount of diversity, that is, for a given set of species, converge to an equilibrium. However, if one wants to understand macroevolutionary changes in the “tempo and mode” (Simpson 1944) of evolution, one needs to consider not only how diversity changes over evolutionary time but also how such changes in diversity affect the nature of evolutionary dynamics (Pennell and Harmon 2013). Indeed, there is evidence from evolution experiments with microbes that the evolutionary dynamics in more diverse communities are qualitatively different from the evolutionary dynamics in less diverse communities (Lawrence et al. 2012). Here we present a theoretical investigation of the question of how diversity affects the complexity of coevolutionary dynamics.

In general, the number of different phenotypes that affect ecological and evolutionary processes is an important quantity. For example, determining the dimensionality of niche space in ecological food webs is a classical problem (Cohen

* Both authors contributed equally to this work.

† Corresponding author; e-mail: doebeli@zoology.ubc.ca.

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1977; Eklöf et al. 2013), and it has recently been shown that including more phenotypic dimensions in models for community assembly has a strong effect on the structure of the emerging food webs (Allhoff et al. 2015). Implicitly, the importance of the dimension of phenotype space is also acknowledged in phylogenetic research through the notion of “adaptive zones” (Simpson 1944; Uyeda et al. 2011). In particular, it is thought that much of the extant diversity has evolved as a consequence of lineages entering new adaptive zones, which can be interpreted from the phenotypic perspective as an increase in the dimension of phenotype space. In general, given the large number of phenotypic properties that determine an individual’s life history and ecology in almost any species, one would expect that ecological interactions are generally determined by many phenotypic properties and that selection pressures emerging from ecological interactions in turn affect many phenotypes simultaneously.

For example, comprehensive modeling of the metabolic network in *Escherichia coli* cells comprises more than 2,000 reactions (Yoon et al. 2012). These reactions are in turn controlled by thousands of genes in a complicated interaction network whose exact workings are largely unknown. Nevertheless, many of the genes contributing to this network of metabolic reactions will be under selection in any given environmental setting, and as a consequence, a large number of phenotypic properties have the potential to undergo evolutionary change. It is generally not known how exactly these phenotypic properties impinge on the birth and death rates of individual organisms and hence what exactly are the ecological selection pressures on these properties. Nevertheless, it seems clear that, in general, many phenotypes will evolve at the same time, that is, that evolution generally takes place in high-dimensional phenotype spaces.

We have recently argued that if evolution takes place in high-dimensional phenotype spaces, then the evolutionary dynamics, that is, the phenotypic change over evolutionary time, can be very complicated, that is, nonstationary and often chaotic (Doebeli and Ispolatov 2014; Ispolatov et al. 2015). In low-dimensional phenotype spaces, nonequilibrium evolutionary dynamics are less likely. However, if a species evolving on a simple attractor gives rise to diversification, the effective dimensionality of the evolving system increases, because the species that emerge from diversification coevolve, driven by both intra- and interspecific ecological interactions. Thus, the total dimensionality of the resulting dynamical system describing multispecies coevolution is the number of species times the dimensionality of the phenotype space in which each species evolves. From our earlier results (Doebeli and Ispolatov 2014; Ispolatov et al. 2015), one could then expect that, as a result of the increase in dimensionality, diversification leads to more

complicated evolutionary dynamics in each of the coevolving species. On the other hand, as a multispecies community becomes more diverse and evolves toward saturation, the available niches tend to get filled, and hence evolutionary change has to become highly coordinated between interacting species and thus constrained, potentially leading to simplified evolutionary dynamics. It is thus unclear how the nature of the evolutionary dynamics changes as the pattern of diversity changes during community assembly.

We investigate these issues by applying the framework of adaptive dynamics (Geritz et al. 1998) to a general class of competition models. The main question we address is, how does the complexity of long-term coevolutionary dynamics depend on the diversity of the coevolving community? We show that in low-dimensional phenotype spaces, there is a hump-shaped relationship between diversity and the complexity of evolutionary dynamics: in communities with low diversity, coevolutionary dynamics are often simple, that is, stationary in the long-time limit; for intermediate degrees of diversity, nonstationary (complex) coevolutionary dynamics are common, and each of the species in the community evolves on a complicated trajectory in phenotype space; and for high amounts of diversity, coevolutionary dynamics become simple again, that is, stationary. In particular, as communities reach diversity saturation through adaptive diversification (Doebeli 2011), coevolutionary dynamics change from complex to simple.

Our results are relevant for a number of issues concerning patterns of macroevolution. For example, the results suggest that, during processes of adaptive radiation (Schluter 2000; Gavrillets and Losos 2009; Seehausen 2015), evolutionary dynamics are more complicated early in the radiation than late in the radiation, a pattern that corresponds to the “early-burst” perspective of macroevolution that has attracted much attention in recent years (Gavrillets and Losos 2009; Harmon et al. 2010; Slater and Pennell 2013). Our results also show that the level at which diversity saturates depends on the dimensionality of phenotype space, with higher dimensions allowing for more diversity. This observation is in accordance with data from radiations in fishes (Seehausen 2015) and points to the possibility of a microevolutionary mechanism for the “blunderbuss theory” of temporal patterns of macroevolutionary changes and diversification (Uyeda et al. 2011): if evolution operates on the dimension of phenotype space on a very slow timescale, then on shorter timescales diversity may saturate and thereby generate relatively stationary evolutionary dynamics, whereas on longer timescales the dimension of phenotype space may increase, for example, because of gene duplications, thus generating a new burst of nonequilibrium (co)evolutionary dynamics until the diversity reaches a new saturation level. Such patterns of intermittent bursts have recently been found in the phylogenies of birds and echinoids (Brusatte et al. 2014; Hopkins and Smith 2015),

and the bursts have been attributed to the evolution of flight capabilities and novel feeding techniques, respectively, both of which can be interpreted as an increase in the dimensionality of the relevant phenotype space. This perspective may also shed light on the question of whether diversity saturates or not (Harmon and Harrison 2015; Rabosky and Hurlbert 2015): diversity may saturate for a given dimension of phenotype space, but evolutionary innovation in the form of new phenotypic dimensions may intermittently generate room for additional bouts of evolutionary diversification.

Methods

Single-Cluster Adaptive Dynamics

As in Doebeli and Ispolatov (2014), we study a general class of models for frequency-dependent competition in which ecological interactions are determined by d -dimensional phenotypes, where $d \geq 1$. For simplicity, we consider homogeneous systems, so no spatial coordinates are included. The ecological interactions are described by a competition kernel $\alpha(\mathbf{x}, \mathbf{y})$ and a carrying capacity $K(\mathbf{x})$, where $\mathbf{x}, \mathbf{y} \in \mathbb{R}^d$ are the d -dimensional continuous phenotypes of competing individuals. The competition kernel α measures the competitive impact that an individual of phenotype \mathbf{x} has on an individual of phenotype \mathbf{y} , and we assume that $\alpha(\mathbf{x}, \mathbf{x}) = 1$ for all \mathbf{x} . Assuming logistic ecological dynamics, $K(\mathbf{x})$ is then the equilibrium density of a population that is monomorphic for phenotype \mathbf{x} . The adaptive dynamics of the phenotype \mathbf{x} is a system of differential equations for $d\mathbf{x}/dt$. To derive the adaptive dynamics, one defines the invasion fitness $f(\mathbf{x}, \mathbf{y})$ as the per capita growth rate of a rare mutant phenotype \mathbf{y} in the monomorphic resident \mathbf{x} population that is at its ecological equilibrium $K(\mathbf{x})$:

$$f(\mathbf{x}, \mathbf{y}) = 1 - \frac{\alpha(\mathbf{x}, \mathbf{y})K(\mathbf{x})}{K(\mathbf{y})}. \quad (1)$$

The expression for the invasion fitness reflects the fact that the growth rate of the mutant \mathbf{y} is negatively affected by the effective density experienced by the mutant, $\alpha(\mathbf{x}, \mathbf{y})K(\mathbf{x})$, discounted by the carrying capacity $K(\mathbf{y})$ of the mutant (see Doebeli 2011 and Doebeli and Ispolatov 2014 for more details). Note that $f(\mathbf{x}, \mathbf{x}) = 0$ for all \mathbf{x} . The invasion fitness $f(\mathbf{x}, \mathbf{y})$ gives rise to the selection gradients in the $i = 1, \dots, d$ phenotypic components:

$$s_i(\mathbf{x}) \equiv \left. \frac{\partial f(\mathbf{x}, \mathbf{y})}{\partial y_i} \right|_{\mathbf{y}=\mathbf{x}} = - \left. \frac{\partial \alpha(\mathbf{x}, \mathbf{y})}{\partial y_i} \right|_{\mathbf{y}=\mathbf{x}} + \frac{\partial K(\mathbf{x})}{\partial x_i} \frac{1}{K(\mathbf{x})}, \quad (2)$$

The selection gradients, in turn, define the adaptive dynamics as a system of differential equations on phenotype space \mathbb{R}^d , which is given by

$$\frac{d\mathbf{x}}{dt} = \mathbf{M}(\mathbf{x}) \cdot \mathbf{s}(\mathbf{x}). \quad (3)$$

Here $\mathbf{s}(\mathbf{x})$ is the column vector $(s_1(\mathbf{x}), \dots, s_d(\mathbf{x}))$, and $\mathbf{M}(\mathbf{x})$ is the mutational variance-covariance matrix. In this matrix, the diagonal elements contain information about the size and rate of mutations in each of the phenotypic dimensions, whereas the off-diagonal elements contain information about the covariance between mutations in two different phenotypic dimensions. This matrix essentially captures the “evolvability” of a population, and it generally depends on the current resident phenotype \mathbf{x} and influences the speed and direction of evolution. For simplicity, we assume here that this matrix is the identity matrix. For more details on the derivation of the adaptive dynamics (eq. [3]), we refer to a large body of primary literature (e.g., Dieckmann and Law 1996; Geritz et al. 1998; Dieckmann 2004; Leimar 2009; Doebeli 2011). We note that the adaptive dynamics (eq. [3]) can be derived analytically as a large-population limit of an underlying stochastic, individual-based model that is again defined on the basis of the competition kernel $\alpha(\mathbf{x}, \mathbf{y})$ and the carrying capacity $K(\mathbf{x})$ (Dieckmann and Law 1996; Champagnat et al. 2006, 2008).

Specifically, here we consider a class of systems that are defined by competition kernels of the form

$$\alpha(\mathbf{x}, \mathbf{y}) = \exp \left[\sum_{i,j=1}^d b_{ij}(x_i - y_i)x_j - \sum_{i=1}^d \frac{(x_i - y_i)^2}{2\sigma_i^2} \right]. \quad (4)$$

Here the coefficients b_{ij} in the first sum on the right-hand side are arbitrary and correspond to the simplest form of a generic, nonsymmetric competition kernel that can generate nonstationary evolutionary dynamics. It can be interpreted as the lowest-order (nontrivial) term from a Taylor expansion of an unknown nonsymmetric competition function. Adaptive dynamics of asymmetric competition have been studied quite extensively (e.g., Law et al. 1997; Kisdi 1999; Doebeli 2011) and are necessary to generate single-species nonequilibrium dynamics in high-dimensional phenotype spaces (Doebeli and Ispolatov 2013, 2014). The second sum on the right-hand side represents “Gaussian competition,” according to which the competitive impact between individuals increases with phenotypic similarity between the competing individuals. The parameters σ_i measure how fast the effect of competition declines as phenotypic distance in the i -component increases. For the carrying capacity we assume

$$K(\mathbf{x}) = \exp \left(- \frac{\sum_{i=1}^d x_i^4}{4} \right). \quad (5)$$

This implies that the carrying capacity imposes an element of stabilizing selection for the phenotype $\mathbf{x} = 0$, at which the carrying capacity is maximal. Thus, the frequency-dependent component of selection is generated by the competition kernel, whereas the frequency-independent compo-

ment of selection is due to the carrying capacity. With these assumptions, the adaptive dynamics (eq. [3]) become

$$\frac{dx_i}{dt} = \sum_{j=1}^d b_{ij}x_j - x_i^3, \quad (6)$$

where $i = 1, \dots, d$. We note that the terms $-x_i^3$ in equation (6) are due to the carrying capacity and serve to contain the trajectories of equation (6) in a bounded domain of phenotype space. Also, the Gaussian part of the competition kernel does not affect the adaptive dynamics of monomorphic populations, that is, the σ_i do not appear in equation (6), because the Gaussian part always has a maximum at the current resident, and hence the corresponding first derivative in the selection gradient (eq. [2]) is 0.

The system of ordinary differential equations (eq. [6]) describes the trajectory of an evolving monomorphic population in phenotype space \mathbb{R}^d . In Doebeli and Ispolatov (2014), we have shown that, for general competition kernels α , such trajectories can be very complicated, particularly when the dimension d is large. With complex evolutionary dynamics, trajectories can be quasiperiodic or chaotic and typically visit many different regions of phenotype space over evolutionary time. When d is low, the dynamics tend to be simpler and often converge to an equilibrium attractor. We can assess the likelihood of equilibrium dynamics for a given dimension d by choosing the d^2 coefficients b_{ij} in equation (6) randomly and independently, for example, from a normal distribution with mean 0 and variance 1, solving the resulting adaptive dynamics (eq. [6]) and checking whether they converge to an equilibrium. If this is done repeatedly, we can approximate the probability of equilibrium dynamics as the fraction of runs that converge to an equilibrium. For $d = 1$, the probability of equilibrium dynamics is, of course, 1, and for $d = 2, 3$, and 4, the resulting probabilities of equilibrium dynamics are approximately 85%, 81%, and 74%, respectively. These are the dimensions that we primarily use in the analysis presented below, but we note that the probability of equilibrium dynamics goes to 0 for large d (Edelman 1997; Doebeli and Ispolatov 2014).

Multicenter Adaptive Dynamics

Here we are interested in the question of how diversification and subsequent coexistence of species (also called “phenotypic clusters” or simply “clusters” throughout the text) affect the evolutionary dynamics. While the Gaussian term in the competition kernel (eq. [4]) does not affect the adaptive dynamics of single monomorphic populations, this term is crucial for determining whether evolutionary diversification occurs. For one-dimensional phenotype spaces ($d = 1$), this is very well known and is encapsulated in the concept of evolutionary branching (Metz et al. 1992; Geritz et al.

1998; Doebeli 2011). An evolutionary branching point is an equilibrium point of equation (6) that is both an attractor for the adaptive dynamics and a fitness minimum. The reason that such points exist in the competition models considered here is precisely that the Gaussian term does not affect the adaptive dynamics but does affect the curvature of the fitness landscape, that is, the second derivative of the invasion fitness (eq. [1]). In particular, small enough σ_i in the Gaussian term will make any equilibrium point a fitness minimum and hence will give rise to evolutionary diversification. Evolutionary branching in scalar traits has been described in a plethora of different models (for an overview, we refer to É. Kisdi’s website at the Department of Mathematics and Statistics at the University of Helsinki, <http://www.mv.helsinki.fi/home/kisdi/addyn.htm>). In high-dimensional phenotype spaces, equilibrium points of equation (6) can also be fitness minima along some directions in phenotype space. For this to happen, the Hessian matrix of second derivatives of the invasion fitness (eq. [1]), evaluated at the equilibrium, must have positive eigenvalues. Indeed, in higher-dimensional phenotype spaces the conditions for the existence of positive eigenvalues of this Hessian matrix, and hence for diversification, generally become less stringent (Doebeli and Ispolatov 2010; Débarre et al. 2014; Svardal et al. 2014).

Importantly, evolutionary diversification can also occur from nonequilibrium adaptive dynamics trajectories (Ito and Dieckmann 2014; Ispolatov et al. 2016). If the adaptive dynamics (eq. [6]) exhibit nonequilibrium dynamics, the crucial quantity determining whether diversification occurs is again the Hessian matrix of second derivatives of the invasion fitness (eq. [1]), but now restricted to the subspace of phenotype space that is orthogonal to the selection gradient (Ispolatov et al. 2016). Essentially, diversification can occur in orthogonal directions in which this Hessian has positive curvature and hence the invasion fitness has a minimum. Because the population is still evolving along the selection gradient, elucidating the exact conditions for diversification requires a careful analysis (Ito and Dieckmann 2014). In the present context, the implication of these results is that, just as with equilibrium adaptive dynamics, diversification can occur along nonequilibrium trajectories of equation (6) if the σ_i in the Gaussian term of the competition kernel are small enough, that is, if the frequency dependence generated by Gaussian competition is strong enough (Ispolatov et al. 2016).

To investigate the process of diversification and the subsequent coevolutionary dynamics, we extend the adaptive dynamics (eq. [6]) to several coexisting phenotypic clusters as follows. We assume that an evolving community consists of m monomorphic populations, each given by a phenotype \mathbf{x}_r , where $r = 1, \dots, m$, with phenotypic components x_{ri} , where $i = 1, \dots, d$ (where d is the dimension of phenotype space). Let N_r be the population density of cluster \mathbf{x}_r . Then

the ecological dynamics of the m clusters are given by the system of logistic differential equations

$$\frac{dN_r(t)}{dt} = N_r(t) \left(1 - \frac{\sum_{s=1}^m \alpha(\mathbf{x}_s, \mathbf{x}_r) N_s(t)}{K(\mathbf{x}_r)} \right), \quad (7)$$

where $r = 1, \dots, m$. Let N_r^* ($r = 1, \dots, m$) denote the equilibrium of system (7) (more generally, for the purposes of deriving the adaptive dynamics, the quantities N_r^* are suitable time averages of population densities over the ecological attractor of eq. [7]; however, our extensive numerical simulations indicated that eq. [7] always converges to an equilibrium). Making the traditional adaptive dynamics assumption that ecological dynamics occur on a faster time-scale than evolutionary dynamics, we calculate the invasion fitness function in cluster r on the basis of the densities N_r^* of the various clusters:

$$f(\mathbf{x}_1, \dots, \mathbf{x}_m, \mathbf{x}_r') = 1 - \frac{\sum_{s=1}^m \alpha(\mathbf{x}_s, \mathbf{x}_r') N_s^*}{K(\mathbf{x}_r')}. \quad (8)$$

Here $\mathbf{x}_1, \dots, \mathbf{x}_m$ describe the phenotypic state of the resident population, and \mathbf{x}_r' denotes the mutant trait in cluster r , where $r = 1, \dots, m$.

Taking the derivative of equation (8) with respect to \mathbf{x}_r' and evaluating it at the resident, $\mathbf{x}_r' = \mathbf{x}_r$, yields the components of the selection gradient \mathbf{s}_r for the cluster r as

$$s_{ri} = \sum_s N_s^* \left(-\frac{1}{K(\mathbf{x}_r)} \frac{\partial \alpha(\mathbf{x}_s, \mathbf{x}_r')}{\partial x_{ri}'} \Big|_{\mathbf{x}_r' = \mathbf{x}_r} + \frac{\alpha(\mathbf{x}_s, \mathbf{x}_r)}{K^2(\mathbf{x}_r)} \frac{\partial K(\mathbf{x}_r)}{\partial x_{ri}} \right), \quad (9)$$

where $i = 1, \dots, d$. For coevolutionary adaptive dynamics, one has to take into account that the rate of mutations in each evolving phenotypic cluster is proportional to the current population size of that cluster (Dieckmann and Law 1996) and hence that the speed of evolution is influenced by the population size. In the single-cluster system, such consideration only rescales time without affecting the geometry of the trajectory and thus is usually ignored. However, in the multicluster system, instead of assuming that the mutational process is described by the identity matrix, as in equation (3), we now assume that, in each cluster r , the mutational variance-covariance matrix \mathbf{M}_r is a diagonal matrix with entries proportional to N_r^* . Assuming that the constant of proportionality is the same in all phenotypic directions, this constant can be set to 1 by rescaling time. This generates the following $d \cdot m$ differential equations describing the adaptive dynamics in the coevolving community:

$$\frac{dx_{ri}}{dt} = N_r^* s_{ri}, \quad (10)$$

where $i = 1, \dots, d$ and $r = 1, \dots, m$. For the multicluster adaptive dynamics, equations (8)–(10) replace their single-

cluster analogs (eqq. [1]–[3]). It is important to note not only that the Gaussian part of the competition kernel α affects whether diversification occurs but also, in contrast to the adaptive dynamics of single monomorphic populations, that the Gaussian term will indeed affect the coevolutionary adaptive dynamics (eq. [10]) of the phenotypic clusters that coexist after diversification has occurred, because it affects both the ecological dynamics (eq. [7]) and the selection gradient (eq. [9]).

Numerical Procedure

To study diversification and subsequent multicluster adaptive dynamics, we implemented the following iterative numerical scenario:

Step 1. Each simulation run is initiated with a randomly generated $d \times d$ matrix of the coefficients b_{ij} for the competition kernel (eq. [4]). The coefficients are drawn from a Gaussian distribution with 0 mean and $d^{-1/2}$ variance. As explained in Doebeli and Ispolatov (2014), this is done to keep the sum of the d terms $\sum_{j=1}^d b_{ij} x_j$ in equation (6) of order x_i , that is, independent of d . Then a certain number of clusters, given by a parameter m_0 , each with population size of order 1, are randomly placed near the phenotype 0, that is, near the maximum of the carrying capacity.

Step 2. For a given set of phenotypic clusters, the ecological dynamics (eq. [7]) are used to solve the population dynamics of all clusters. The system of differential equations is integrated via a fourth-order Runge-Kutta algorithm for $\sim 10^3$ time steps of duration $dt \sim 10^{-2}$ to ensure convergence to the equilibrium (or, in case there is no such convergence, to ensure a correct calculation of the time average of the various population densities). If the population density of a given cluster falls below the threshold $N_{\min} \sim 10^{-8}$, the cluster is eliminated from the system. During the ecological dynamics, the evolutionary dynamics are frozen and evolutionary time does not advance.

Step 3. After calculation of the N_r^* , where $r = 1, \dots, m$ and m is the current number of clusters, the adaptive dynamics of the phenotypes of the clusters are advanced via equations (9) and (10) with a fourth-order Runge-Kutta algorithm with a typical time step $d\tau \sim 10^{-2}$, by which the evolutionary time is advanced as well. After this evolutionary time step, the ecological dynamics are recalculated, potentially preceded by the following step 4, which is performed only if the corresponding time condition is satisfied.

Step 4. The level of diversity, that is, the number of clusters in the system, is controlled as follows. Each τ_c time units the distances between clusters are assessed. If the dis-

tance between two or more clusters is below a threshold, $\Delta x \sim 10^{-3}$, these clusters are merged, preserving the total population size of the merged clusters and the position of their center of mass. Immediately after this comparison step, the total number of clusters is compared to the target number of clusters, which is given by a system parameter m_{\max} . If the current number of clusters is below m_{\max} , a new cluster is created by randomly picking an existing cluster, splitting it in half and separating the two new clusters in a random direction in phenotype space by the distance of the merging threshold, Δx .

Step 5. In our simulations, we take measurements at regular time intervals (τ_m , ranging from ~ 1 to 10 time units). One of the main quantities of interest is the average per capita evolutionary speed v in the evolving community, which is the average of the norms of the vectors of trait variation (evolution) rates in each cluster, weighted by the cluster population size, computed as

$$v = \frac{\sum_{r=1}^m N_r \sqrt{\sum_{i=1}^d (dx_{ri}/dt)^2}}{\sum_{r=1}^m N_r} \quad (11)$$

This quantity is a strong indicator of the nature of the evolutionary dynamics of the coevolving system. In particular, our very extensive numerical simulations indicate that when the average speed falls below 10^{-2} , then the system eventually exhibits equilibrium evolutionary dynamics. In contrast, when the average evolutionary speed remains high, the coevolving system tends to exhibit complicated, non-equilibrium dynamics, with the majority of the clusters exhibiting large fluctuations in phenotype space over evolutionary time. An example of such nonequilibrium coevolution is given in the next section. Other measurements include the position and population size of all clusters in the system and the number of “distinct” clusters separated by a “visible” distance $\Delta X = 0.1$. These measurements can also be averaged over time.

For any given simulation run initiated by step 1 above, steps 2–5 were repeated iteratively until a specified final simulation time was reached or until evolution came to a halt, which by our definition occurs when the average evolutionary speed falls below a threshold, $v < 10^{-4}$. Our general approach consisted of simulating many different systems according to the above scheme and then computing statistical characteristics such as the fraction of runs that resulted in nonequilibrium dynamics or the average evolutionary speed as a function of the level of diversity (see “Results”).

One crucial feature of our algorithm is the periodic generation of new clusters in step 4, which mimics diversification events, that is, evolutionary branching. Diversification

is thus modeled by simply adding new phenotypic clusters at certain points in time and close to existing clusters. This mimics the sympatric split of an ancestral lineage. Sympatric diversification is a theoretically robust phenomenon (Doebeli 2011), and our procedure represents a shortcut for this phenomenon necessitated by computational feasibility. If such splitting is not feasible, given the current ecological circumstance, the new cluster will not diverge phenotypically from the ancestor and hence will be merged again with the ancestor (see below). Alternatively, newly generated clusters may become extinct ecologically. In either case, speciation is not successful. Thus, in our models it is the ecological circumstances that determine whether speciation can occur or not, but the process of speciation itself (i.e., the splitting) is performed in a simplified manner. If speciation is successful and the newly generated clusters diverge and persist ecologically, then diversity has increased (unless other clusters become extinct). We note that, by construction, the maximal level of diversity in a given simulation run, that is, the number of different clusters, cannot exceed the parameter m_{\max} . Therefore, this parameter allows us to control the level of diversity in a given simulation.

There are in principle other, less artificial ways to model diversification. In particular, stochastic, individual-based models and partial differential equation models (Champagnat et al. 2006, 2008) have been used to describe the evolutionary dynamics of phenotype distributions. In such models, diversification is an emergent property that is reflected in the formation of new modes in the evolving phenotype distributions. While these techniques are very useful in general, they are currently not computationally feasible for the statistical approach that we employed here, which requires systematic simulation of many different systems. Also, they would not allow for control of the level of diversity, as the number of phenotypic modes would simply be an emergent property of the evolving system. Nevertheless, we have used these alternative techniques to illustrate the robustness of salient results for particular examples. A more detailed description of these techniques is given in the appendix, available online. Another alternative would be to assume that new clusters (species) are assigned phenotypes that are chosen randomly in phenotype space, rather than close to an existing cluster. This could correspond to immigration of new species into an existing community. However, we would not expect this to affect our main results, because with complicated evolutionary dynamics, the initial phenotypic position of a given cluster becomes irrelevant after some time.

Finally, we note that the merging of clusters (species) is done solely for computational reasons and has no biological meaning (apart from designating organisms that are closely related and phenotypically very close as belonging to the same species). Merging of clusters occurs only shortly after

a new cluster is seeded close to an existing one and only if the new cluster does not diverge from the existing one (i.e., only if the ecological conditions for diversification are not satisfied). If divergence is successful, the clusters will never again get close enough to other clusters to be merged, because of the repelling force of frequency-dependent competition. Thus, the only function of merging is to prevent the number of clusters from artificially becoming very large.

Results

The parameter that controls the level of diversity in our simulations is m_{\max} , which is the maximal number of different phenotypic clusters allowed to be present at any point in time in an evolving community (see step 4 in “Numerical Procedure”). Our first result is obtained by allowing this parameter to be very large, so that we can estimate the number of clusters that eventually coexist by simply running the simulations for a long time and recording the number of clusters at which the diversity equilibrates. We denote by $M_{\sigma,d}$ the equilibrium number of clusters for a given phenotypic dimension d and strength of the Gaussian component σ in the competition kernel (eq. [4]). We found that the equilibrium level of diversity increases exponentially with the dimension d of phenotype space and decreases with σ (fig. 1). Here and below we assume for simplicity that the σ_i are the same in all phenotypic directions, $\sigma_i = \sigma$ for $i = 1, \dots, d$. In the appendix, we indicate scaling relationships that hold for $M_{\sigma,d}$ as functions of the parameters σ and d . In general, diversity is maintained only if $\sigma \lesssim 1$, which is roughly the scale of the phenotypic range set by the carrying capacity (eq. [5]). Only if $\sigma \lesssim 1$ does the equilibrium level of diversity increase exponentially with increasing dimension of phenotype space.

Our main results are now obtained on the basis of the observation that, if the parameter m_{\max} is fixed at a value smaller than $M_{\sigma,d}$ for a given d and σ , the community will typically evolve to a diversity level m_{cluster} of approximately m_{\max} . That is, if the diversity is constrained to be below the maximal level of diversity possible for a given set of parameters, then the diversity will typically evolve to the value set by the constraint. Note that this is an “average” statement about many simulation runs, that is, many different choices of the coefficients b_{ij} and stochastic realizations of cluster splitting. While some simulation runs will result in a diversity that is lower than m_{\max} (which may reflect an intrinsic state of the system for the given set of coefficients or a long-living metastable state that has not yet reached its full diversity), most runs will evolve to the level of diversity that is prescribed by this parameter.

This allows us to then assess, for a given level of diversity, the nature of the coevolutionary dynamics that unfold in communities with that level of diversity. Two paradigmatic

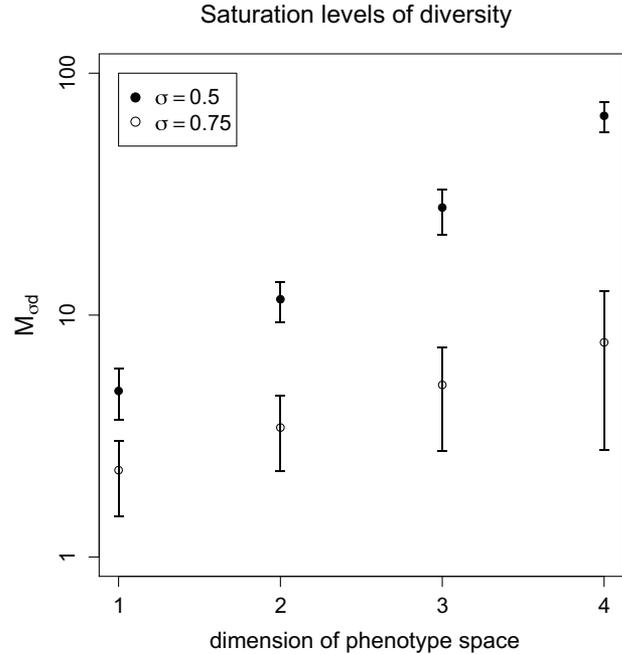


Figure 1: Exponential growth of the equilibrium number of clusters $M_{\sigma,d}$ versus phenotype dimension d for $\sigma = 0.5$ and 0.75 . For each dimension, $M_{\sigma,d}$ was determined numerically as an average from several hundred simulation runs in which the parameter m_{\max} was set sufficiently high.

examples are shown in figure 2. We first set the level of diversity $m_{\max} = 12$, which is far below the saturation level $M_{\sigma,d}$ for the given system. Starting from very few clusters, the diversity quickly evolves to the level set by m_{\max} , and the coexisting clusters then exhibit complicated, nonstationary evolutionary dynamics, with all clusters undergoing sustained and irregular fluctuations in phenotype space (fig. 2, left). This type of complicated dynamics is characterized by average evolutionary speeds $v > 10^{-2}$. In the same system, but now with a value of m_{\max} that lies above the saturation level $M_{\sigma,d}$, the diversity evolves to the saturation level, at which the community consists of ~ 30 coexisting phenotypic clusters (fig. 2, right). In this saturated state, the average evolutionary speed is much lower than 10^{-2} and the community exhibits much more stationary coevolutionary dynamics (that would eventually converge to a coevolutionary equilibrium). Moreover, the saturated community exhibits a characteristic pattern of overdispersion in phenotype space due to competitive repulsion caused by the Gaussian component of the competition kernel (see also fig. A1, available online).

To obtain a more systematic characterization of the coevolutionary dynamics as a function of the diversity of the evolving community, we ran, for a given dimension of phenotype space d and strength of competition σ , 100

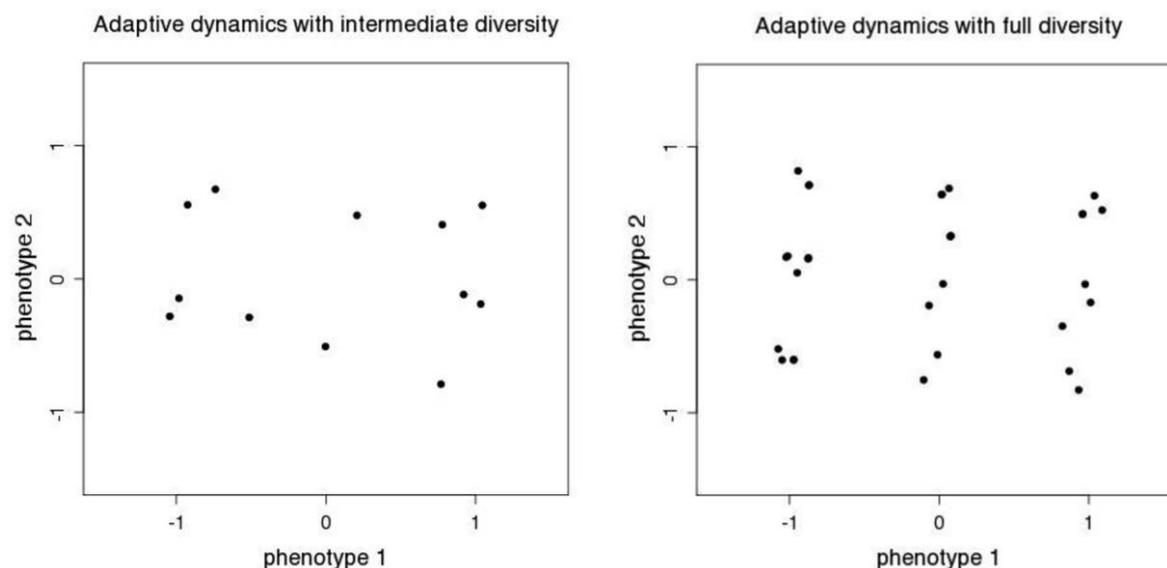


Figure 2: Snapshots of the end of the videos illustrating coevolutionary dynamics with different levels of diversity. The videos can be found in the appendix (videos A1, A2), available online. In the snapshots and videos, phenotypes are projected onto the first two phenotypic dimensions. For these examples, the competition kernel was defined by the coefficients b_{ij} given in the appendix, and $d = 3$ and $\sigma = 0.5$. In the left-hand panel, $m_{\max} = 12$, and in the right-hand panel $m_{\max} = 40$. The saturation level of diversity for this system is $M_{\sigma,d} \sim 30$. With fewer (12) clusters, that is, less diversity (*left*), the coevolutionary dynamics are nonstationary and the clusters undergo sustained and irregular fluctuations in phenotype space, as is apparent in the corresponding video. At the diversity saturation level (*right*), the evolutionary dynamics slow down and become almost stationary. The systems were run for 400 time units, and the average evolutionary speed at the end of the simulation runs was $\nu = 3.49 \times 10^{-2}$ in the first case, and $\nu = 5.54 \times 10^{-4}$ in the second case.

simulations with randomly chosen coefficients b_{ij} for each $m_{\max} = 1, \dots, M$, where M is some number that is larger than the saturation level of diversity $M_{\sigma,d}$. For each run, we recorded the average per capita evolutionary speed ν and the number of phenotypic clusters, that is, the level of diversity, present at the end of 1,000 evolutionary time units (averaged over the last 4 time units). We classified the dynamics into equilibrium dynamics if the average speed ν was less than 10^{-2} and nonequilibrium dynamics otherwise. As mentioned above, this was based on individual inspection of many simulation that ran longer than 1,000 time units, which showed that the threshold 10^{-2} is a very good indicator of whether the coevolutionary system eventually equilibrates.

Our main results are shown in figures 3 and 4. The general pattern is that the probability of nonequilibrium dynamics increases as diversity increases from single-cluster communities to communities with a few clusters (fig. 3). For intermediate diversity, the fraction of nonequilibrium dynamics remains high. For communities with high diversity, the fraction of nonequilibrium dynamics starts to decrease, and almost all communities with a diversity close to the saturation level $M_{\sigma,d}$ exhibit equilibrium coevolutionary dynamics. To illustrate these trends, we give a more detailed account of the average velocities ν defined in equation

(11) in the coevolving communities (fig. 4). It shows that there is an exponential decrease in the average speed as the diversity increases and that there is a substantial fraction of low-diversity communities that exhibit equilibrium dynamics. The exact shape of these patterns depends on d and σ (figs. 3, 4), but whenever diversification is possible, the overall trend is that nonequilibrium dynamics are most likely at intermediate levels of diversity and that high levels of diversity tend to generate equilibrium coevolutionary dynamics.

The patterns shown in figures 3 and 4 are based on many different simulated communities with different levels of diversity. However, similar patterns can be observed in simulations of single communities as they evolve from low to high diversity, that is, as they undergo an adaptive radiation. Such a radiation, starting from a single phenotypic cluster, is shown in figure 5A. Over time, the evolving community becomes more diverse as a result of adaptive diversification, and as a consequence the nature of the coevolutionary dynamics of the community changes. In the example shown in figure 5A, the coevolutionary dynamics are fast for low to intermediate levels of diversity and then slow down as the community acquires more and more species, until eventually the community reaches a coevolutionary equilibrium at

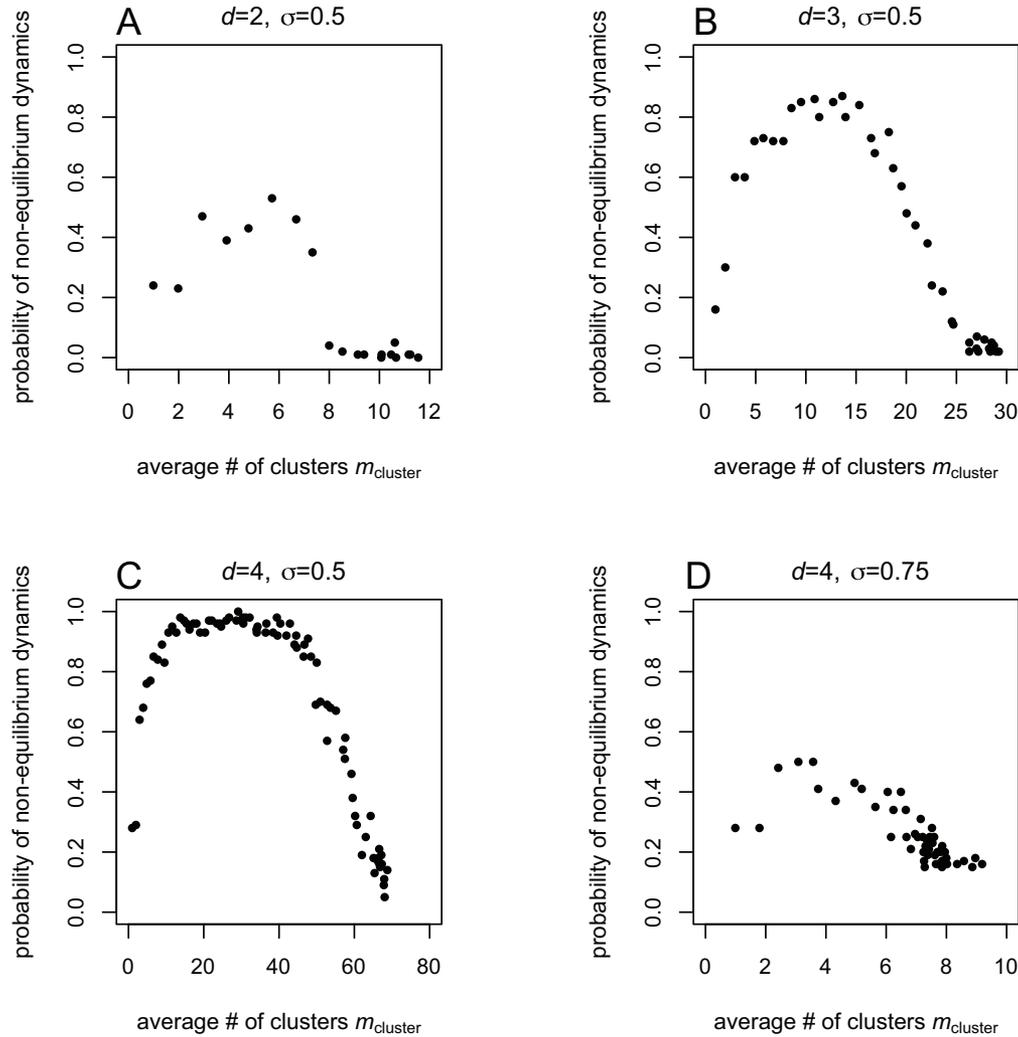


Figure 3: Probability of nonequilibrium evolutionary dynamics versus the number of distinct clusters m_{cluster} present in the evolving community for $d = 2, \sigma = 0.5$ (A); $d = 3, \sigma = 0.5$ (B); $d = 4, \sigma = 0.5$ (C); $d = 4, \sigma = 0.75$ (D). For each panel, the results were obtained by running, for each integer value of m_{max} in an interval $[1, M]$ with $M > M_{\sigma,d}$ (the diversity threshold), 100 simulations with different random initial conditions and sets of b_{ij} coefficients for 1,000 time units. For each of these simulation runs, we recorded the final evolutionary speed and the final number of clusters. We then determined what fraction of the 100 simulation runs for a given m_{max} had an evolutionary speed above the equilibrium dynamics threshold of 10^{-2} and plotted that fraction against the average of the final cluster number calculated from the 100 simulations.

the diversity saturation level. Again, the slowdown of the evolutionary speed during an adaptive radiation appears to occur exponentially with an increase in diversity. This can also be seen by running a given community defined by a given set of coefficients b_{ij} for different values of the parameter m_{max} , determining the level of diversity possible in the evolving community. The evolutionary speed exponentially decreases with the diversity given by m_{max} (fig. 5B). We currently do not have a mechanistic explanation for the exponential decay in evolutionary rates with increasing diversity. It is informative to watch the process of diversification and subsequent evolu-

tionary slowdown unfold dynamically, for which we refer to the videos in the appendix. To verify that the observed dynamical pattern is not an artifact of the adaptive dynamics approximation, we performed individual-based and partial differential equation simulations of the same system. The movies in videos A3–A5 in the appendix (videos A1–A5 available online), corresponding to the scenario used for figures 2 (right) and 5A, confirm that all three methods produce qualitatively similar evolutionary pictures. The detailed descriptions of the individual-based and partial differential equation methods are given in the appendix.

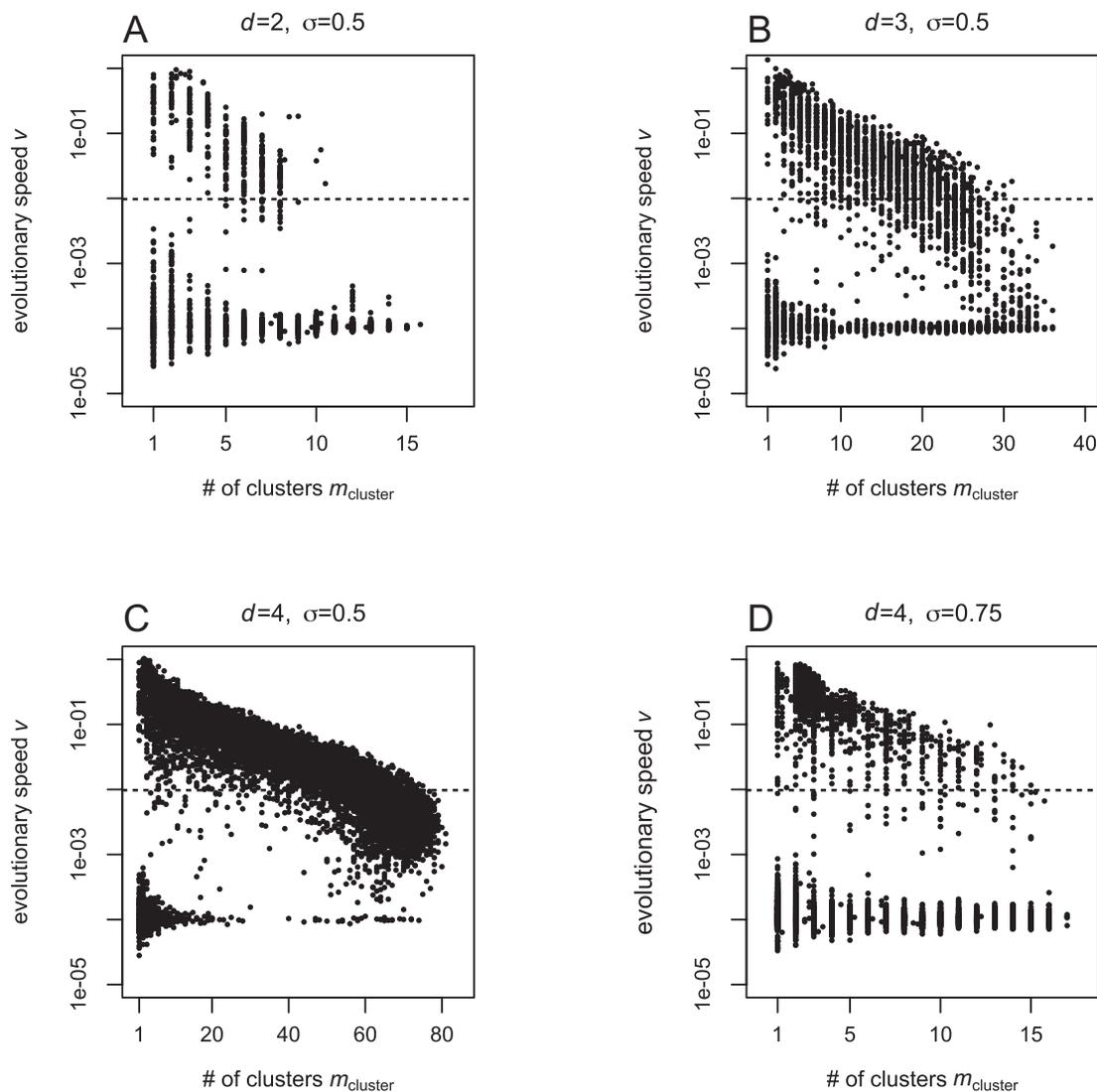


Figure 4: Per capita evolutionary speed v versus the number of distinct clusters m_{cluster} for $d = 2, \sigma = 0.5$ (A); $d = 3, \sigma = 0.5$ (B); $d = 4, \sigma = 0.5$ (C); and $d = 4, \sigma = 0.75$ (D). The points at the bottom of the plots at around $v = 10^{-4}$ correspond to systems that reached the steady state before the final time of the simulation and the corresponding runs were stopped when the average velocity was below this threshold. For each panel, the results were obtained from the same simulations runs that were used for figure 3, by plotting the final evolutionary speed versus the final cluster number for each simulation run, averaged over a short period at the end of the run (so that the number of clusters is not necessarily an integer). The dashed horizontal lines indicate the threshold for equilibrium evolutionary dynamics (see text).

Another interesting, although perhaps not so surprising, observation for single adaptive radiations concerns the rate of accumulation of new species in the evolving ecosystem. Figure 5C shows the number of species as a function of time during the adaptive radiation scenario used for figure 5A, illustrating that the rate of diversification is highest at the beginning of the radiation and then slows as the community evolves toward the diversity saturation level. The details of these dynamics depend on system parameters and, in particular, on the rate at which new species are introduced into the system, but the qualitative behavior of diversification

rates, which are initially high and then slow down, is common to all adaptive radiations generated by our models.

Discussion

We investigated the expected long-term evolutionary dynamics resulting from competition for resources in models for gradual evolution in high-dimensional phenotype spaces. In reality, most organisms have many different phenotypic properties that impinge on their ecological interactions in generally complicated ways, and here we assumed that multi-

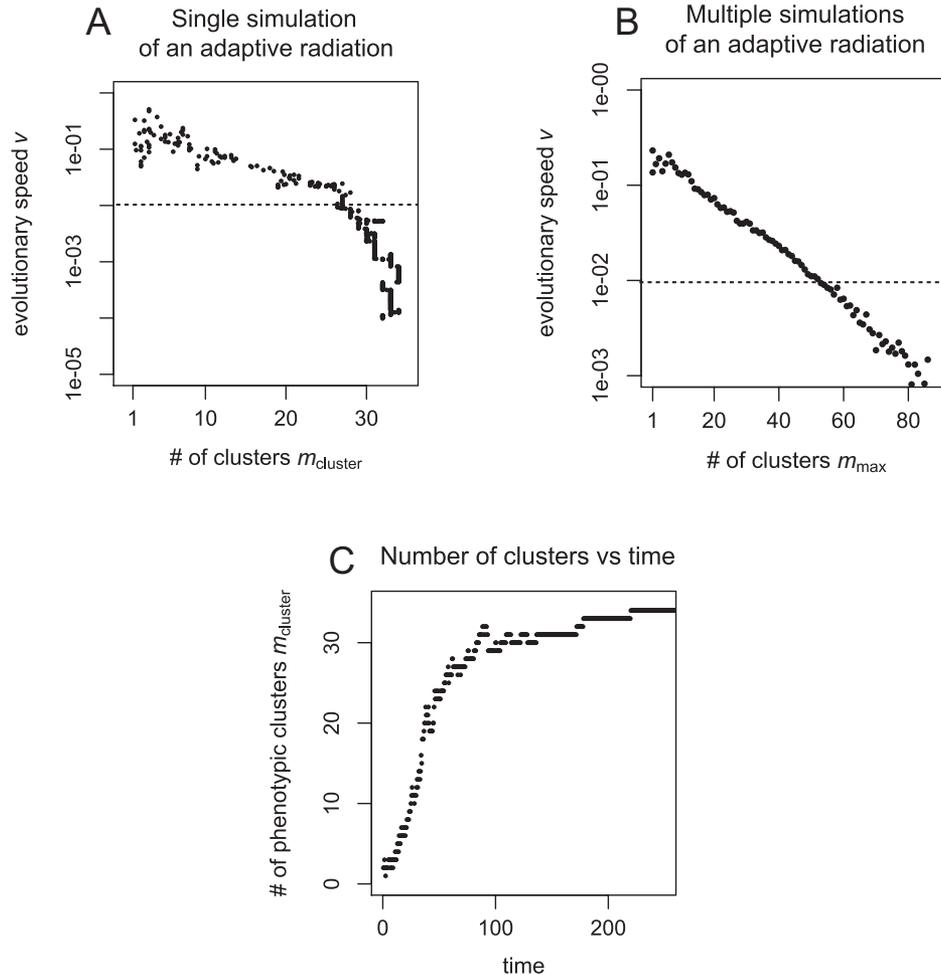


Figure 5: A, Per capita evolutionary speed v versus the number of distinct clusters $m_{cluster}$ present in the evolving community for a single simulation run, corresponding to a single adaptive radiation. This is the same scenario as that shown in figure 2 (right): the competition kernel was defined by the coefficients b_{ij} given in the appendix, available online, and $d = 3$, $\sigma = 0.5$, and $m_{max} = 40$. The system exhibits equilibrium single-cluster dynamics. Note the initial increase of the speed v as the system evolves from a single-cluster to a few-clusters regime, followed by a steady exponential decay of v until the evolving community is diverse enough to exhibit equilibrium dynamics. B, Per capita evolutionary speed v as a function of the maximum number of clusters m_{max} for a choice of coefficients b_{ij} for $d = 4$ and $\sigma = 0.5$. The coefficients b_{ij} are given in the appendix. The system exhibits nonequilibrium single-cluster dynamics. For each m_{max} the system was run for 1,000 time units and the per capita velocity v was averaged over the last 200 time steps. Note that while the case shown in B corresponds to a single choice of the coefficient matrix b_{ij} , separate simulations were run for different values of m_{max} . This is in contrast to A, which shows data from a single simulation run (with a large m_{max}). The dashed horizontal lines indicate the threshold for equilibrium evolutionary dynamics (see text). C, Number of distinct phenotypic clusters (species) as a function of time for the scenario shown in A. The rate of diversification is high at the beginning of the radiation and slows down as the evolving community reaches saturation diversity.

dimensional phenotypes determine logistic ecological dynamics through the competition kernel and the carrying capacity. We then used coevolutionary adaptive dynamics algorithms to extend the ecological dynamics to macroevolutionary timescales, and we used a statistical approach to capture general properties of the ensuing evolutionary dynamics.

If the negative frequency dependence generated by the competition kernel is strong enough, competition results

in repeated adaptive diversification and hence in communities of coevolving phenotypic species. By randomly choosing many different competition kernels, we showed that the complexity of the coevolutionary dynamics in such communities is expected to be highest for intermediate levels of phenotypic diversity. In particular, as the evolving communities increase in diversity toward the saturation level, that is, the maximal number of different species that can coexist, the evolutionary dynamics become simpler and com-

munities at the saturation level are expected to exhibit a co-evolutionary equilibrium. We also showed that the diversity saturation level increases exponentially with the dimension of phenotype space.

Our interpretation of these findings is that in low-dimensional phenotype spaces such as the ones considered here, evolutionary dynamics of single species are expected to converge to an equilibrium (Doebeli and Ispolatov 2014). However, as diversity increases, the different phenotypic clusters will “push” each other around evolutionarily as a result of frequency-dependent competition. This occurs mostly because of the repulsive nature of pairwise interaction induced by the Gaussian term in the competition kernel (eq. [4]): clusters that move farther apart decrease competition felt from each other. For example, the splitting of a cluster sitting at an attractive fixed point of the adaptive dynamics creates two offspring clusters that may diverge from each other as a result of competitive repulsion. As long as diversity is not very high, that is, as long as there is enough available niche or unoccupied phenotype space, this typically results in nonequilibrium coevolutionary dynamics if the competition kernel is not symmetric, thus leading to an increase in evolutionary complexity with phenotypic diversity. As the diversity keeps increasing toward saturation levels, which for each phenotypic dimension is determined roughly by the ratio of the widths of the carrying capacity and the competition kernel (see video A2 in the appendix), the available niche space gets filled, so that the evolving clusters “have nowhere to go” evolutionarily. An analogy with gas-liquid-solid phase transitions may illustrate this in the following way. As in the dynamics of molecules, the adaptive dynamics of phenotypic clusters contain a pairwise-repulsive term, which originates from the Gaussian term in the competition kernel. A few-clusters regime qualitatively corresponds to the gas phase, when the range of the repulsive interaction is significantly less than the typical distance between clusters. As the number and thus density of clusters increases, the repulsive interaction becomes more relevant, constraining the individual motion of clusters and resulting in a liquid-like behavior, where clusters are predominantly localized and occasionally hop to a new location. Finally, the maximum cluster density creates a crystal-like structure, albeit not necessarily entirely symmetric because of the randomly generated b_{ij} terms in the adaptive dynamics. The motion of individual clusters is heavily constrained by its neighbors via mutual repulsion, while the collective motion of an ensemble of clusters is limited by the carrying-capacity function. Thus, phenotypic saturation leads to a state in which the coevolving clusters are strongly constrained evolutionarily by the other clusters in the community, and hence to coevolutionary equilibrium dynamics.

Some empirical support for an initial increase in the complexity of evolutionary dynamics with the number of

species in an ecosystem comes from the laboratory evolution experiments of Lawrence et al. (2012), who showed that the speed of adaptation to novel environments is higher in bacterial species that are part of microbial communities with a small number of competitors than when evolving in monoculture. However, our results are seemingly in contrast to previous theoretical results about the effect of diversity on evolutionary dynamics (de Mazancourt et al. 2008; Johansson 2008). These authors essentially argued that while a single species is free to evolve in response to changes in the environment, evolution of the same species is more constrained in a community of competitors, in which other species are more likely to evolutionarily occupy new niches. Hence, diversity is expected to slow down evolution. However, these models describe evolution only in one-dimensional phenotypes and may thus miss the complexity arising in higher-dimensional spaces. Moreover, even in higher-dimensional spaces, the arguments for evolutionary slowdown presented in de Mazancourt et al. (2008) and Johansson (2008) essentially correspond to our observation of a slowdown when diversity reaches saturation, at which point evolutionary change in each species is indeed constrained because of the competing species occupying all available niches. Our approach also has to be distinguished from approaches based primarily on ecological dynamics, as in Shtilerman et al. (2015). In these approaches, emerging ecological communities are also modeled by periodically adding new species, but there is no underlying phenotype space that would determine competitive interactions. Instead, every time a new species is added, its interaction coefficients with the already existing species are chosen according to a specific, randomized procedure. This leads to interesting results, such as saturating levels of diversity after initially fast and fluctuating increases from low levels of diversity. However, since there is no underlying phenotype space, this approach does not reveal the evolutionary dynamics of continuous phenotypes, and, in particular, it does not yield any information about the effects of the dimension of phenotype space on the evolutionary dynamics or on the amount of diversity at saturation.

There has been much interest in recent years in determining the effects of phylogenetic relationships on the functioning of ecosystems (e.g., Ives and Godfray 2006; Cadotte et al. 2013; Nuismer and Harmon 2015). The intuitive notion is that phylogenetic information has predictive power for ecological interactions if recently diverged species are more likely to interact than those that diverged long ago. More specifically, Nuismer and Harmon (2015) have argued that phylogenetic information is most likely to be relevant for ecosystem dynamics if ecological interactions are based on phenotypic matching, so that species with more similar trait values are more likely to interact strongly. Our models have a component of phenotypic matching due to the Gaussian part of

the competition kernel, but they also have a strong component of different types of interactions due to the “random” part of the competition kernel given by the coefficients b_{ij} . As we have shown, it is this non-Gaussian part of the competition kernel that causes the complicated coevolutionary dynamics, and it is this complexity, in turn, that makes phylogenetic signal largely irrelevant in our models.

A full phylogenetic analysis of the macroevolutionary dynamics generated by our models is beyond the scope of this work, but we can provide some basic insights based on the complicated evolutionary dynamics in phenotype space that the different phenotypic clusters (species) perform when there is an intermediate number of clusters in the coevolving community. An example of this is shown in video A1. Here, after an initial phase of diversification, the community contains 12 coevolving clusters. These clusters move on a complicated evolutionary trajectory, with each cluster undergoing large evolutionary changes without further diversification. No matter what the phylogenetic relationship between these clusters (as given by their emergence from the single initial cluster), it is clear that because of the large evolutionary fluctuations in phenotype space of each cluster (species), there will be no consistent correlation between phylogenetic relationship and phenotypic distance. Even if there were such a correlation (positive or negative) at a particular point in time, it would change over time because of the large evolutionary fluctuations of each cluster over time. This is illustrated in figure A1A, which shows that no persistent correlation pattern between phylogenetic and phenotypic distance should be expected in communities with an intermediate amount of diversity. In particular, recently diverged species are not more likely to interact than those diverged less recently, because the evolving community has a short “phenotypic memory” as a result of complicated evolutionary dynamics.

However, when further diversification is allowed, so that the system reaches its saturation level of diversity, not only does the coevolving community become more diverse but the evolutionary dynamics slow down, leading to ever smaller phenotypic fluctuations. In particular, new clusters emerging toward the end of the assembly of the evolutionarily stable community will stay phenotypically closer to their phylogenetically most closely related clusters, that is, to their parent or sister species. Therefore, in the last phase of community assembly, a positive correlation between phylogenetic and phenotypic distance can be expected to build up, at least to some extent. This is illustrated in figure A1B. Thus, weak phylogenetic signals are expected to develop toward the end of community assembly.

Regarding adaptive radiations, two observations emerge from our models. The first concerns the classical notion that rates of diversification should decline over the course of a radiation (Schluter 2000; Gavrillets and Losos 2009), a

pattern that seems to have good empirical support (Schluter 2000; McPeck 2008; Rabosky and Lovette 2008; Gavrillets and Losos 2009). Our models confirm this pattern of declining rates of diversification (fig. 5C). The second observation is that rates of evolution should generally slow down with an increase in diversity. This should be true not only when different ecosystems are compared (figs. 3, 4) but also during an adaptive radiation in a single evolving community (fig. 5). Thus, we would expect the evolutionary dynamics to be faster and more complicated early in an adaptive radiation and to slow down and eventually equilibrate late in the radiation. This corresponds to the so-called early-burst model of macroevolution (Gavrillets and Losos 2009; Harmon et al. 2010) in the context of adaptive radiations. This model predicts that when lineages enter novel “adaptive zones” (Simpson 1944), such as novel ecological niches, evolutionary rates in the lineage should be fast initially and then slow down as the adaptive zone gets filled with diverse phenotypes. Harmon et al. (2010) found little evidence for the early-burst model when analyzing a large set of data from many different clades. Nevertheless, these authors noted that younger clades have higher rates of evolution than older clades, which points to the fact that evolutionary rates may slow down with clade age. Moreover, few clades in their data set correspond to the type of very fast adaptive radiation envisaged and observed in our models, and they did not consider high-dimensional phenotypes. Finally, Harmon et al. (2010) note that groups with a larger proportion of sympatric species early in their history would be more likely to exhibit an early-burst pattern. In our models, adaptive radiations occur in complete sympatry and indeed produce the early-burst pattern.

According to Slater and Pennell (2013), the jury on early-burst models is still out, and in fact substantial evidence for this model has accumulated in recent years. For example, Cooper and Purvis (2010) reported an early burst in body size evolution in mammals, Weir and Mursleen (2013) observed an early-burst pattern in the evolution of bill shape during adaptive radiation in seabirds, Gonzalez-Voyer and Kolm (2011) and Arbour and López-Fernández (2013) reported early-burst patterns in morphological and functional evolution in cichlids, and Benson et al. (2014) described patterns of early bursts in the evolution of dinosaur morphology.

Uyeda et al. (2011) have incorporated the early-burst concept into a macroevolutionary perspective in which, over very long evolutionary timescales, rare but substantial phenotypic bursts alternate with more stationary periods of bounded phenotypic fluctuations, somewhat reminiscent of the concept of punctuated equilibrium (Gould and Eldredge 1977) when applied to rates of phenotypic evolution (Pennell et al. 2014). We think that the models presented here could provide a microevolutionary basis for such a perspective if they are extended by considering evolutionary

change in the dimension of the phenotype space that determines ecological interactions. Such an extended theory would have three timescales: a short, ecological timescale, an intermediate timescale at which coevolution and diversifications take place in a given phenotype space, and a long timescale at which the number of phenotypic components increases (or decreases). Our hypothesis would then be that in such systems, periods of bounded evolutionary fluctuations near diversity saturation levels for a given dimension of phenotype space would alternate with bursts of rapid evolutionary change, brought about by an evolutionary increase in phenotypic dimensions and the subsequent increase in diversity and acceleration in evolutionary rates until a new saturation level is reached. The resulting long-term evolutionary dynamics would thus show periods of relative phenotypic stasis alternating with periods of fast evolution. This picture would fit very well with the “blunderbuss” pattern envisaged in Uyeda et al. (2011). These authors proposed that the intermittent bursts in evolutionary rates are caused by lineages encountering novel “adaptive zones” (Simpson 1944). Novel adaptive zones would correspond to the opening up of new habitats or new resources, which would in turn correspond to new phenotypes that determine use of the novel adaptive zone. Alternatively, novel adaptive zones could also be generated by the emergence of novel sets of regulatory mechanisms, allowing novel uses of already existing habitats and resources (as, e.g., when a trade-off constraint is overcome through gene duplication). In either case, novel adaptive zones would correspond to an increase in the dimensionality of ecologically important phenotypes.

It is interesting to note that such intermittent-burst patterns have in fact been observed in phylogenetic data and that they seem to be connected to novel, ecologically important phenotypes. Hopkins and Smith (2015) have shown that evolutionary rates in echinoids reveal at least two instances of rapidly accelerating and subsequently declining evolutionary rates, that is, two intermittent bursts. Moreover, these bursts appear to be associated with the evolution of novel feeding strategies (Hopkins and Smith 2015; Slater 2015). Also, Brusatte et al. (2014) have shown that an evolutionary burst occurs in the dinosaur-bird transition, and it is tempting to conjecture that this burst was caused by the increase in phenotype dimensionality due to the proliferation of flight capabilities.

There is also good empirical support for our finding that the level at which diversity saturates increases with the dimension of phenotype space. Seehausen (2015) has argued that, essentially, the high number of different ecologically relevant traits is the basis for the spectacular radiations of cichlids in African lakes. In conjunction with ecological opportunity, genetic and phenotypic flexibility, which appears to be at least in part due to gene duplications, has allowed

this group of fish to reach a much higher diversity than other groups, such as cichlids in rivers or whitefish in arctic lakes, in which fewer phenotypes appear to be ecologically relevant (Hudson et al. 2011; Seehausen 2015). In this context, we note that incorporating the evolution of the dimension of phenotype space may also shed light on the ongoing debate about whether diversity saturates over evolutionary time (Harmon and Harrison 2015; Rabosky and Hurlbert 2015). It seems that the answer could be “yes and no”: diversity saturates in the intermediate term for a given dimension of phenotype space but does not saturate in the long term if the dimension of phenotype space increases over long evolutionary timescales, thus generating recurrent increases in saturation levels.

Our study has a number of limitations that should be addressed in future research. It is currently impractical to perform the statistical analysis presented here for phenotype spaces with dimensions higher than 4 because of computational limitations. Our results indicate that the diversity saturation level, that is, the maximal number of coexisting phenotypic clusters, increases rapidly with the dimension d of phenotype space, which makes simulations of communities at saturation levels unfeasible. Nevertheless, we expect the salient result that coevolutionary dynamics slow down as communities reach the saturation level to be true in any dimension as long as the Gaussian component of competition in equation (4) affects all phenotypic directions. Also, in our approach we have assumed that the phenotypes determining competitive interactions are the same for intra- and interspecific competition. This may be a fair assumption for closely related species, such as those coevolving in an adaptive radiation. However, for competition in more general ecosystems, it may also be relevant to assume that from a total set of d phenotypes, different subsets determine competition within a species and competition with various other species. In addition, to describe general ecosystems and food webs, it will be important to include not just competitive interactions but also predator-prey and mutualistic interactions, each again determined by potentially high-dimensional phenotypes. Also, throughout we have assumed a simple unimodal form of the carrying capacity to represent the external environment. More complicated forms of the carrying capacity, and hence of the external fitness landscape, will likely generate even richer patterns of coevolutionary dynamics and diversification. Finally, we have assumed throughout that evolving populations are well mixed, and it will be interesting to see how the results generalize to spatially structured ecosystems. All these extensions remain to be developed.

We are, of course, aware of the fact that we did not include genetic mixing due to sexual reproduction in our models, and our method of describing diversification by simply adding new phenotypic clusters, although fairly standard, does not take into account the actual process of

speciation. In sexual populations, adaptive diversification due to disruptive selection, as envisioned here, requires assortative mating, and the conditions for the evolution of various types of assortative mating, as well as for the likelihood of speciation once assortment is present, have been studied extensively (e.g., Doebeli 2011). A general, if crude, conclusion from this work is that when there is enough disruptive selection for diversification to occur in asexual models, then it is likely that adaptive speciation also occurs in the corresponding sexual models, although factors such as the strength of assortment, population size, and linkage disequilibrium may become important. It would, in principle, be possible to incorporate sexual reproduction into the models presented here, for example, along the lines of Gascuel et al. (2015). Our previous results (Doebeli and Ispolatov 2010; Ispolatov et al. 2016) indicate that adaptive diversification is generally more likely in high-dimensional phenotype spaces, and we think that the present models serve well as a first approximation to study adaptive diversification and coevolutionary dynamics in evolving communities.

Ultimately, the applicability and relevance of our models for understanding macroevolutionary patterns in nature depend in part on being able to determine evolutionary rates of high-dimensional phenotypes from phylogenetic data, which appears to be a difficult problem (Adams 2013, 2014; Denton and Adams 2015). Nevertheless, overall we think that our approach of incorporating microevolutionary processes based on ecological interactions in high-dimensional phenotype spaces into statistical models for macroevolutionary dynamics has the potential to shed new light on a number of fundamental conceptual questions in evolutionary biology.

Acknowledgments

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**Online Appendix A: Diversity and long-term
coevolutionary dynamics in high-dimensional
phenotype spaces**

1 Videos for Figure 2 in main text

Video A 1: Video illustrating coevolutionary dynamics for an intermediate level of diversity, corresponding to the still image in Fig. 2a in the main text. Phenotypes are projected onto the first two phenotypic dimensions. For this examples, the competition kernel was defined by the coefficients $\{b_{ij}\}$ given in Section 5, and $d = 3$, $\sigma = 0.5$, and $m_{max} = 12$. The saturation level of diversity for this system is $M_{\sigma,d} \sim 30$. With intermediate diversity (12 clusters), the coevolutionary dynamics is non-stationary and the clusters undergo sustained and irregular fluctuations in phenotype space. The system was run for 400 time units, and the average evolutionary speed at the end of the simulation runs was $v = 3.49 \times 10^{-2}$.

Video A 2: Video illustrating coevolutionary dynamics with a high level of diversity, corresponding to the still image in Fig. 2b in the main text. Phenotypes are projected onto the first two phenotypic dimensions. For this examples, the competition kernel was defined by the coefficients $\{b_{ij}\}$ given in Section 5, and $d = 3$, $\sigma = 0.5$, and $m_{max} = 40$. The saturation level of diversity for this system is $M_{\sigma,d} \sim 30$. As the evolving community reaches the diversity saturation level, the evolutionary dynamics slows down and becomes almost stationary. The systems were run for 400 time units, and the average evolutionary speed at the end of the simulation runs was $v = 5.54 \times 10^{-4}$.

2 Correlation between phylogenetic and phenotypic distance

For each pair of clusters (species) in an evolving community we define the phylogenetic distance between them, Pg , as the number of links in the path between them on the phylogenetic tree. To measure this distance, we add the following scheme to our evolutionary algorithm:

- The system is initialized with a single cluster.
- Each cluster splitting event produces two offspring separated by the distance 2. The distance between an offspring and all its existing neighbours is incremented by one.
- When two recently split cluster that failed to diverge are merged, the distance between the newly produced common cluster and each of its neighbours is calculated as the minimum of the distances of the two merged clusters minus one. This reflects the observation that merging events only happen with newly split clusters.

As a result, at any given time we know phylogenetic distances between all pairs of clusters currently present in the system. To quantify the relation between the phenotypic and phylogenetic similarity, we compute the correlation C between phylogenetic and phenotypic distance as follows:

$$C = \frac{\langle [Pg - \langle Pg \rangle][X - \langle X \rangle] \rangle}{\sigma_{Pg}\sigma_X}, \quad (1)$$

where Pg and X are phylogenetic and phenotypic distances between clusters, $\langle \dots \rangle$ define the average over all pairs of clusters present in the system and σ_{Pg} and σ_X are the standard deviations of distances.

The above scheme allows us to track the correlation between phylogenetic and phenotypic distance over time, as illustrated in Figure A1. Fig. A1A shows the time dependence of C for the simulation shown in Video A1, and in Fig. A1B shows the time dependence of C for the simulation shown in Video A2. During the early phase of community assembly the correlation C rapidly decays due to complicated coevolutionary dynamics of the emerging clusters. When the diversity of the coevolving community is kept intermediate (by setting the parameter m_{max} to intermediate values, as in Video A1), the correlation between phylogenetic and phenotypic distance itself undergoes fluctuations around 0 (Fig. A1A). This is because the clusters in the community with intermediate diversity undergo large phenotypic fluctuations while their phylogenetic relationship is constant, because no further diversification (or extinction) occurs. However, when the diversity is allowed to reach saturation levels (by setting m_{max} to a large value, as in Video A2), a positive correlation between phylogenetic and phenotypic distance develops in the final stages of community assembly, i.e., as the coevolving community reaches the saturation diversity and hence undergoes much smaller phenotypic fluctuations (Fig. A1B). Note that the correlation is still close to 0 during the early stages of community assembly, but some correlation remains at the end due to clusters emerging in the last phase of community assembly, which tend to stay phenotypically closer to their sister species because evolutionary dynamics become slow and stable.

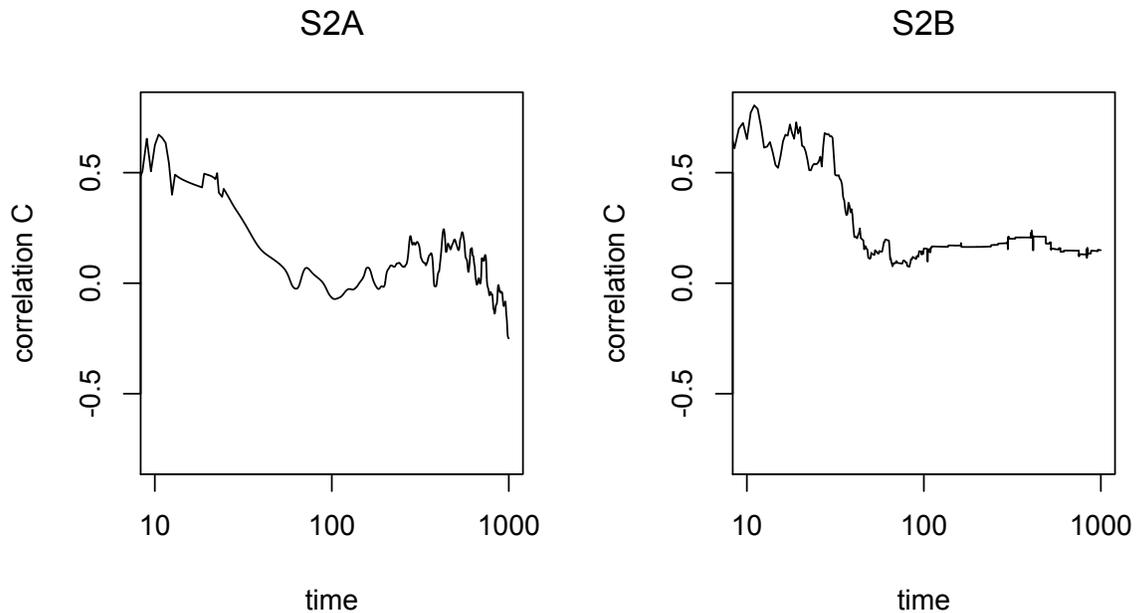


Figure A 1: Correlation between phylogenetic and phenotypic distance in an evolving community as a function of time. A1A) corresponds to the simulation run shown in Video A1, and A1B) corresponds to the simulation run shown in Video A2. In A1A), the correlation decays and then fluctuates around 0, reflecting the fact that diversity is kept intermediate, leading to complicated evolutionary dynamics, during which individual clusters (species) undergo large phenotypic fluctuations, while the phylogenetic relationships are constant once the maximal allowed level of diversity is reached. In A1B), the community is allowed to reach the saturation level of diversity. During the later stages of community assembly, the evolutionary dynamics slow down and become more stable, allowing for a positive correlation between phylogenetic and phenotypic distance in the youngest species of the community. Note the the log scale on the time axis.

3 Individual-based simulations

Individual-based realizations of the model were based on the Gillespie algorithm (Gillespie, 1976) and consisted of the following steps:

1. The system is initialized by creating a set of $K_0 \sim 10^3 - 10^4$ individuals with phenotypes $\mathbf{x}_k \in \mathbf{R}^d$ localized around the initial position \mathbf{x}_0 with a small random spread $|\mathbf{x}_k - \mathbf{x}_0| \sim 10^{-3}$.
2. Each individual k has a constant reproduction rate $\rho_k = 1$ and a death rate $\delta_k = \sum_{l \neq k} A(\mathbf{x}_l, \mathbf{x}_k) / [K(\mathbf{x}_k)]$, as defined by the logistic ecological dynamics.
3. The total update rate is given by the sum of all individual rates, $U = \sum_k (\rho_k + \delta_k)$.
4. The running time t is incremented by a random number Δt drawn from the exponential distribution $P(\Delta t) = U \exp(-\Delta t U)$.
5. A particular birth or death event is randomly chosen with probability equal to the rate of this event divided by the total update rate U . If a reproduction event is chosen, the phenotype of an offspring is offset from the parental phenotype by a small mutation randomly drawn from a uniform distribution with amplitude $\mu = 10^{-3} - 10^{-2}$.
6. The individual death rates δ_k and the total update rate U are updated to take into account the addition or removal of an individual.
7. Steps 4-6 are repeated until t reaches a specified end time.

The movie in Video A4 shows the dynamics of the individual-based model corresponding to the adaptive dynamics simulation shown in Video A3, which is the same as the scenario used for Video A2 (note that the movie in Video A3 runs for $t = 1200$ time units, whereas the movie in Video A2 runs for $t = 400$ time units).

Video A 3: Video of an adaptive radiation. This is the scenario from which data for Figure 5A was extracted. The corresponding coefficients b_{ij} for the competition kernel are given in Section 5. Phenotypes are projected onto the first two phenotypic dimensions and shown for three different simulation methods that give qualitatively similar results. This video shows the same multicluster adaptive dynamics simulation as in Video A2, but for a longer time span, $t = 0, \dots, 1200$.

Video A 4: Video of an adaptive radiation. Phenotypes are projected onto the first two phenotypic dimensions and shown for three different simulation methods that give qualitatively similar results. The video shows an individual-based simulation corresponding to the system shown in Video A3. Details of the individual-based simulations are described in Section 3 of the Online Appendix).

Video A 5: Video of an adaptive radiation. The video shows a numerical simulation of the partial differential equation model explain ins Section 4 of the Online Appendix. The scenario shown corresponds to the ones shown in Videos A3 and A4. Shown are the projections of the positions of the maxima of the phenotype distribution in the evolving community. Note that the completely deterministic partial differential equation model is symmetric in that it is invariant under the coordinate inversion $x \rightarrow -x$. Therefore, the solution illustrated in the video is symmetric as well.

4 Partial differential equation models

A deterministic large-population limit of the individual-based model is obtained as the partial differential equation (PDE)

$$\frac{\partial N(\mathbf{x}, t)}{\partial t} = N(\mathbf{x}, t) \left(\frac{1 - \int \alpha(\mathbf{y}, \mathbf{x}) N(\mathbf{y}, t) d\mathbf{y}}{K(\mathbf{x})} \right) + D \sum_{i=1}^d \frac{\partial^2 N(\mathbf{x}, t)}{\partial x_i^2}, \quad (2)$$

where $N(\mathbf{x}, t)$ is the population distribution at time t (Champagnat et al., 2006). The second term of the right hand side is a diffusion term that describes mutations, with the diffusion coefficient typically set to $D \sim 10^{-4} - 10^{-3}$. Local maxima of the solution $N(\mathbf{x}, t)$ can be interpreted as positions of the centers of the phenotypic clusters. Their dynamics are shown in Video A5. For any given scenario, the corresponding adaptive dynamics solution can be used to determine the single- or few-cluster trajectory, and hence to approximately determine the region occupied by the system in phenotype space over time. Note that the deterministic PDE model is invariant with regard to the coordinate change $\mathbf{x} \rightarrow -\mathbf{x}$, and hence its solutions must be symmetric with regard to simultaneous reflection on all coordinate axes. To numerically solve the PDE model (2) we chose a lattice noticeably larger than the corresponding adaptive dynamics attractor. The number of bins B in each dimension of this lattice is strongly constrained by memory limitations: An efficient implementation requires computing and storing an array of $B^d \times B^d$ values of the competition kernel $\alpha(\mathbf{y}_i, \mathbf{x}_j)$ for the pairwise interactions between all pairs of sites i and j . With $B = 25 - 30$ to achieve a reasonable spatial resolution, the memory constraint makes the PDE implementation feasible only for $d = 2, 3$.

The movie in Video A5 shows the dynamics of the partial differential equation model corresponding to the scenarios shown in Videos A3 and A4.

5 Scaling relationship for the diversity at saturation

The number of clusters at the diversity saturation level, $M_{\sigma,d}$, can be estimated to be proportional to the volume of the available phenotype space with the linear dimension L , divided by the volume occupied by each cluster, which has a typical linear size σ :

$$M_{\sigma_a,d} \approx C_\sigma \frac{L^d}{\sigma^d}. \quad (3)$$

Hence, the following scaling relationships hold:

$$M_{\sigma_a,d} = M_{\sigma_b,d} \left(\frac{\sigma_b}{\sigma_a} \right)^d \quad \text{and} \quad M_{\sigma,d_1} = M_{\sigma,d_2}^{d_1/d_2}, \quad (4)$$

where σ_a and σ_b denote different strengths of competition, and C_σ is a constant of order 1 that takes into account the “imperfect packing” occurring when σ and L have similar magnitude. Based on this, the equilibrium level of diversity is expected to increase exponentially with increasing dimension of phenotype space (as illustrated in Figure 1), and with increasing frequency-dependence (i.e., decreasing σ). In general, diversity is only maintained if $\sigma \lesssim 1$, which is roughly the scale of the phenotypic range set by the carrying capacity given by eq. (5) in the main text.

6 Specific sets of coefficients used

The following set of coefficients b_{ij} determining the competition kernel were used for Figures 5A in the main text and for the movies in Videos A1 and A2:

$$\begin{pmatrix} 0.407 & 0.498 & 0.287 \\ -0.199 & -1.102 & -0.305 \\ 1.387 & -0.896 & 0.341 \end{pmatrix} \quad (5)$$

The following set of coefficients b_{ij} determining the competition kernel were used for Figure 5B in the main text:

$$\begin{pmatrix} -1.289 & 0.682 & 0.217 & -0.093 \\ -0.223 & -0.035 & 0.697 & -0.117 \\ -0.563 & 0.434 & -0.953 & -0.198 \\ 0.119 & 0.398 & 0.183 & 0.530 \end{pmatrix} \quad (6)$$

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